

MURI Meeting

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Testing the goodness-of-fit of AR(p) model

$$X_t = \sum_{k=1}^p \phi_k X_{t-k} + Z_t$$

- ▶ $Z_t \stackrel{iid}{\sim} F$
- ▶ Causal
 - ▶ $1 - \sum_{k=1}^p \phi_k z^k = 0$ no roots inside the unit circle

Testing the goodness-of-fit of AR(p) model

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- ▶ Observations $(X_t)_{t=1, \dots, n}$

Testing the goodness-of-fit of AR(p) model

$$X_t = \sum_{k=1}^p \hat{\phi}_k X_{t-k} + \hat{Z}_t$$

- ▶ Observations $(X_t)_{t=1, \dots, n}$
- ▶ Parameter estimates $\hat{\phi}_k$ (e.g., least-square, Yule-Walker, etc.)
- ▶ Fitted residuals (\hat{Z}_t)

Proposal: Test the serial dependence of (\hat{Z}_t)

Measure of Dependence: Generalized Distance Covariance

$$T^2(X, Y; \mu) := \int |\varphi_{X,Y}(s, t) - \varphi_X(s)\varphi_Y(t)|^2 \mu(ds, dt)$$

- ▶ Random variables $X \in \mathbb{R}^p$ and $Y \in \mathbb{R}^q$,

$$X \perp\!\!\!\perp Y \iff \varphi_{X,Y}(s, t) = \varphi_X(s)\varphi_Y(t), \quad \forall s, t$$

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- ▶ Existence:

- ▶ μ finite
- ▶ μ infinite: $\int (1 \wedge |s|^\alpha)(1 \wedge |t|^\alpha) \mu(ds, dt) < \infty$ and $\mathbb{E}[|X|^\alpha + |Y|^\alpha + |XY|^\alpha] < \infty$

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$$\mu(ds, dt) = c|s|^{-2}|t|^{-2}$$

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- ▶ Traditional distance covariance (Székely et al., 07):

$$\mu(ds, dt) = c|s|^{-2}|t|^{-2}$$

- ▶ Distance correlation:

$$R^2(X, Y; \mu) = \frac{T^2(X, Y; \mu)}{\sqrt{T^2(X, X; \mu)T^2(Y, Y; \mu)}}$$

Measure of Dependence: Generalized Distance Covariance

$$T^2(X, Y; \mu) := \int |\varphi_{X,Y}(s, t) - \varphi_X(s)\varphi_Y(t)|^2 \mu(ds, dt)$$

- ▶ Assume $\mu(ds, dt) = \mu_1(ds) \times \mu_2(dt)$ symmetric about the origin
- ▶ Let $\tilde{\nu}(s) = \begin{cases} \int_{\mathbb{R}^d} e^{i\langle s, x \rangle} \nu(dx) & , \quad \nu \text{ finite} \\ \int_{\mathbb{R}^d} (1 - \cos\langle s, x \rangle) \nu(dx) & , \quad \nu \text{ infinite} \end{cases}$

$$\begin{aligned} T(X, Y; \mu) &= \mathbb{E}[\tilde{\mu}_1(X - X') \tilde{\mu}_2(Y - Y')] \\ &\quad + \mathbb{E}[\tilde{\mu}_1(X - X')] \mathbb{E}[\tilde{\mu}_2(Y - Y')] \\ &\quad - 2 \mathbb{E}[\tilde{\mu}_1(X - X') \tilde{\mu}_2(Y - Y'')]. \end{aligned}$$

- ▶ where X', Y', Y'' are iid copies of X, Y, Y respectively

Measure of Dependence: Generalized Distance Covariance

$$T_n^2(\mathbf{X}, \mathbf{Y}; \mu) := \int |\varphi_{\mathbf{X}, \mathbf{Y}}^n(s, t) - \varphi_{\mathbf{X}}^n(s) \varphi_{\mathbf{Y}}^n(t)|^2 \mu(ds, dt)$$

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$$\begin{aligned} T_n(X, Y; \mu) &= \frac{1}{n^2} \sum_{i,j=1}^n \tilde{\mu}_1(X_i - X_j) \tilde{\mu}_2(Y_i - Y_j) \\ &\quad + \frac{1}{n^4} \sum_{i,j,k,l=1}^n \tilde{\mu}_1(X_i - X_j) \tilde{\mu}_2(Y_k - Y_l) \\ &\quad - \frac{2}{n^3} \sum_{i,j,k=1}^n \tilde{\mu}_1(X_i - X_j) \tilde{\mu}_2(Y_i - Y_k). \end{aligned}$$

Measure of Dependence: Generalized Distance Covariance

$$T_n^2(\mathbf{X}, \mathbf{Y}; \mu) := \int |\varphi_{\mathbf{X}, \mathbf{Y}}^n(s, t) - \varphi_{\mathbf{X}}^n(s)\varphi_{\mathbf{Y}}^n(t)|^2 \mu(ds, dt)$$

- ▶ Consistency under **ergodicity**
- ▶ Asymptoticity under **certain α -mixing** condition

Auto-Distance Covariance Function

Let $\mathbf{Z}_1 = (Z_1, \dots, Z_{n-h})$, $\mathbf{Z}_{h+1} = (Z_{h+1}, \dots, Z_n)$, then

$$T_n^2(h) := T_n^2(\mathbf{Z}_1, \mathbf{Z}_{h+1}; \mu)$$

If $Z_t \stackrel{iid}{\sim} F$, then

$$n T_n^2(h) \xrightarrow{d} \int |G_F(s, t)|^2 \mu(ds, dt)$$

- ▶ G_F is a Gaussian field dependent on F

Example: Kilkenny wind speed time series

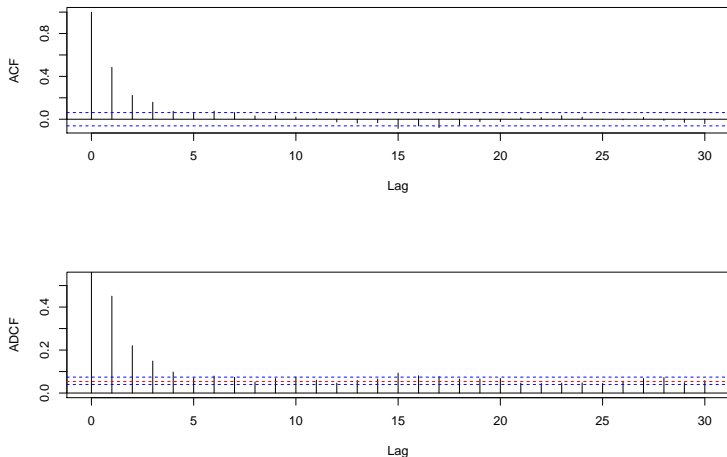


Figure : ACF and auto-distance correlation function (ADCF) of Kilkenny daily wind speed time series from 1/1/61 - 9/27/63

Testing the goodness-of-fit of AR(p) model

$$X_t = \sum_{k=1}^p \hat{\phi}_k X_{t-k} + \hat{Z}_t$$

- ▶ Observations $(X_t)_{t=1, \dots, n}$
- ▶ Parameter estimates $\hat{\phi}_k$
- ▶ Fitted residuals (\hat{Z}_t)

Proposal: Test the serial dependence of (\hat{Z}_t)

Statistic of interest: $\tilde{T}_n^2(h) := T_n^2(\hat{\mathbf{Z}}_1, \hat{\mathbf{Z}}_{h+1}; \mu)$

Auto-distance covariance of AR(p) residuals

$$X_t = \sum_{k=1}^p \hat{\phi}_k X_{t-k} + \hat{Z}_t$$

Statistic of interest: $\tilde{T}_n^2(h) := T_n^2(\hat{\mathbf{Z}}_1, \hat{\mathbf{Z}}_{h+1}; \mu)$

Theorem

- ▶ Assume that

$$\int (s^2 \wedge t^2 \wedge (st)^2) \mu(ds, dt) < \infty$$

- ▶ If $\mathbb{E}Z^2 < \infty$, then

$$n \tilde{T}_n^2(h) \xrightarrow{d} \int |G_F(s, t) + \xi_h(s, t)|^2 \mu(ds, dt)$$

where

$$\xi_h(s, t) = t\varphi_Z(t) \varphi'_Z(s) \Psi_h^T \mathbf{Q}$$

- ▶ $\Psi_h = (\psi_{h-j})_{j=1, \dots, p}$ where ψ_j are the coefficients in the causal representation $X_t = \sum_{j=0}^{\infty} \psi_j Z_{t-j}$
- ▶ \mathbf{Q} is the limit distribution of $\sqrt{n}(\hat{\phi} - \phi)$

Example: Auto-distance covariance of AR(10) residuals: $Z_t \sim N(0, 1)$

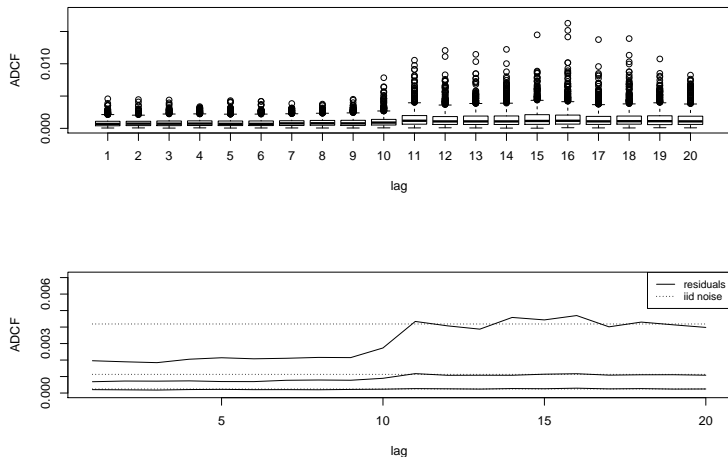


Figure : Left panel: empirical box plots of \tilde{T}_n^2 ; Right panel: empirical 5%, 50%, 95% quantiles of \tilde{T}_n^2 and T_n^2

Example: Auto-distance covariance of AR(10) residuals: $Z_t \sim N(0, 1)$

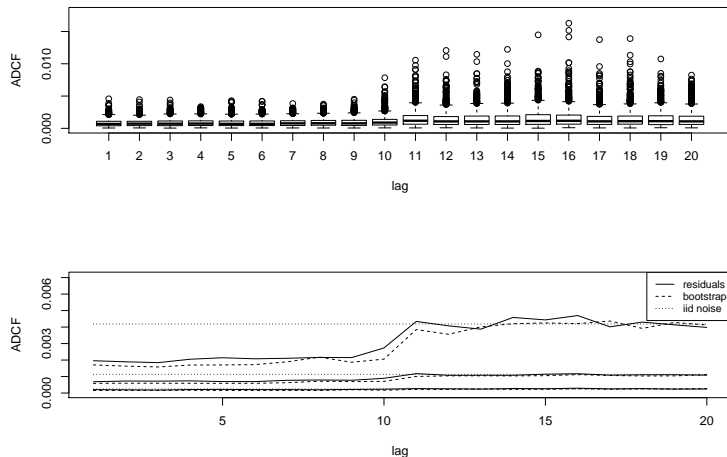


Figure : Left panel: empirical box plots of \tilde{T}_n^2 ; Right panel: empirical 5%, 50%, 95% quantiles of \tilde{T}_n^2 , from simulations and from bootstrapping, and that of T_n^2

Example: Auto-distance covariance of AR(10) residuals

$$X_t = \sum_{k=1}^p \hat{\phi}_k X_{t-k} + \hat{Z}_t$$

Statistic of interest: $\tilde{T}_n^2(h) := T_n^2(\hat{\mathbf{Z}}_1, \hat{\mathbf{Z}}_{h+1}; \mu)$

Theorem

- ▶ Assume that

$$\int (s^2 \wedge t^2 \wedge (st)^2) \mu(ds, dt) < \infty$$

- ▶ If Z is in the domain of attraction of a stable distribution of index $\alpha \in (0, 2)$, then

$$n \tilde{T}_n^2(h) \xrightarrow{d} \int |G_F(s, t)|^2 \mu(ds, dt).$$

Example: Auto-distance covariance of AR(10) residuals: $Z_t \sim t(1.5)$

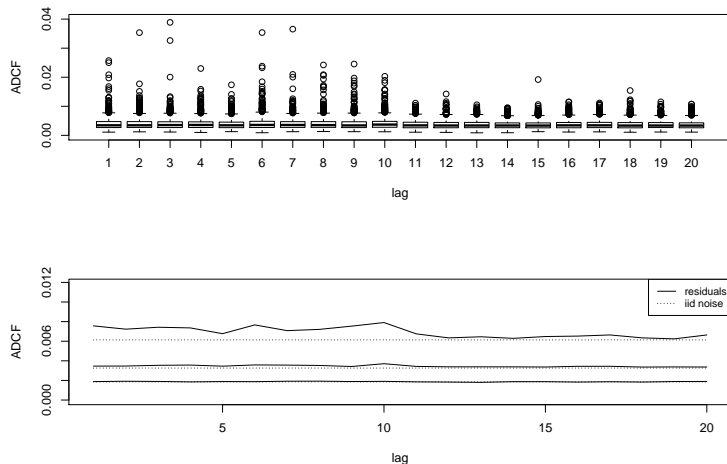


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Example: Kilkenny wind speed time series

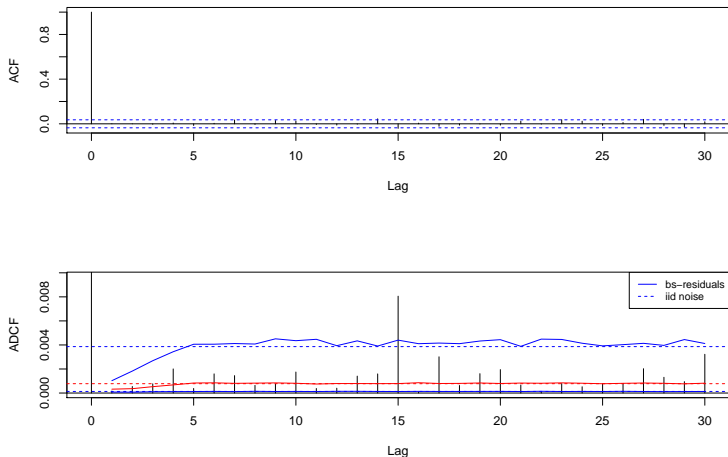


Figure : ACF and auto-distance correlation function (ADCF) of Kilkenny daily wind speed time series after AR(3) fitting.